



A model relating transpiration for Japanese cedar and cypress plantations with stand structure



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ABSTRACT

Previous studies have revealed that changes in forest structure due to management (e.g., thinning, aging, and clearcutting) could affect the forest water balance. However, there are unexplained variability in changes in the annual water balance with changing structure among different sites. This is the case even when analyzing data for specific species/regions. For a more advanced and process-based understanding of changes in the water balance with changing forest structure, we examined transpiration (E) observed using the sap-flux method for 14 Japanese cedar and cypress plantations with various structure (e.g., stem density and diameter) in Japan and surrounding areas and developed a model relating E with structural parameters. We expressed E using the simplified Penman–Monteith equation and modeled canopy conductance (G_c) as a product of reference G_c (G_{cref}) when vapor pressure deficit is 1.0 kPa and functions expressing the responses of G_c to meteorological factors. We determined G_{cref} and parameters of the functions for the sites separately. E observed for the 14 sites was not reproduced well by the model when using mean values of G_{cref} and the parameters among the sites. However, E observed for the sites was reproduced well when using G_{cref} determined for each site and mean values of the parameters of the functions among the sites, similar to the case when using G_{cref} and the parameters of the functions determined for each site. These results suggest that considering variations in G_{cref} among the sites was important to reproduce variations in E , but considering variations in the parameters of the functions was not. Our analysis revealed that G_{cref} linearly related with the sapwood area on a stand scale (A) and that A linearly related with stem density (N) and powers of the mean stem diameter (d_m). Thus, we proposed a model relating E with A (or N and d_m), where G_{cref} was calculated from A (or N and d_m) and the parameters of the functions were assumed to be the mean values among the sites. This model estimates changes in E with changing structure from commonly available data (N and d_m), and therefore helps improve our understanding of the underlying processes of the changes in the water balance for Japanese cedar and cypress plantations.

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1. Introduction

Changes in forest structure due to management (e.g., planting, growth, thinning, aging, and clearcutting) can affect the forest water balance. Numerous studies (Scott and Lesch, 1997; Cornish and Vertessy, 2001) have examined changes in the annual water balance with changing forest structure on the basis of catchment

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water balance data. Summarizing data derived from such studies and analyzing them using linear regression, researchers have identified several important parameters (e.g., the ratio of the cutting area to the total catchment area, annual rainfall, and leaf phenology) determining changes in annual runoff with changing structure (Bosch and Hewlett, 1982; Brown et al., 2005; Komatsu et al., 2011). However, there remains unexplained variability in changes in the annual water balance (Bosch and Hewlett, 1982; Brown et al., 2005). This is the case even for data for a single species within a specific region, although the variability is less pronounced (Adams and Fowler, 2006; Komatsu et al., 2011). This suggests the limitation of analyzing catchment water balance data simply using linear regression without considering underlying processes (e.g., canopy transpiration and interception evaporation). For a more advanced understanding of changes in the annual water balance with changing structure, examining relationships of canopy transpiration and interception for given species with various structure (e.g., stem density and diameter) would be useful (Komatsu et al., 2007d). Variations in stem density and diameter relate with the sapwood area on a stand scale and leaf area index, which could in turn relate with canopy transpiration (Granier et al., 2000a; Ewers et al., 2011). Focusing on specific species would be useful in reducing factors to be considered, because there are factors (e.g., the clumping factor and stem conductivity) that could differ among different species and affect canopy transpiration (Baldochi and Meyers, 1998; Zwieniecki and Holbrook, 1998; Bréda, 2003).

Japanese cedar (*Cryptomeria japonica*) and cypress (*Chamaecyparis obtusa*) are major plantation species in Japan and surrounding areas such as China and Taiwan (Japan Forestry Agency, 2014). Examining the water balance of these plantations is highly important from a practical viewpoint. Forest management (e.g., thinning) has been performed to increase water resources in Japan, although its effectiveness has not been assessed sufficiently (Komatsu et al., 2010). Summarizing data for interception evaporation of Japanese cedar and cypress plantations (Hattori et al., 1982; Tanaka et al., 2005), Komatsu et al. (2007a) found a relationship between stem density and interception evaporation and then developed a model relating interception evaporation with stem density. However, there have been few studies examining the relationship between forest structure and canopy transpiration (E) for Japanese cedar and cypress plantations.

To assess changes in the water balance of Japanese cedar and cypress plantations with changing forest structure, we developed a model relating E for Japanese cedar and cypress plantations with meteorological and structural variables. The model formulates E using the simplified Penman–Monteith equation (McNaughton and Black, 1973; Jarvis and McNaughton, 1986). Canopy conductance (G_c) in the equation was written as a series of functions expressing responses of G_c to meteorological factors (Jarvis, 1976; Lohammer et al., 1980). This study comprises three steps. First, we calculated G_c using E data derived employing the sap-flux method for 14 sites and the inverted form of the simplified Penman–Monteith equation to determine parameters of the model for the sites separately. Second, we assessed the importance of each parameter in determining E on the basis of sensitivity analysis. Third, we examined the relationship of the important parameters with structural parameters. Here, we tried to relate the structural parameters with commonly available data such as stem density and diameter for wide use of the model.

Models estimating E are roughly classified into two groups. One uses many empirical parameters for modeling stomatal/canopy conductance to avoid considering internal hydraulics and keep the model structure simple (Granier et al., 2000a; Komatsu, 2004), while the other considers internal hydraulics (Domec et al., 2012; McDowell et al., 2013). Our model belongs to the former group. Most models of the former group focus on reproducing

E for a specific site (Cienciala et al., 1994a,b; Granier and Bréda, 1996). Several models (Granier et al., 2000a; Komatsu, 2004) are applicable to various sites, similar to our model. Our model differs from these models in that our model focuses on two species, which suggests higher predictability of the model when applied to the species. Furthermore, our model differs from the models in that our model would be tested against E data recorded not only during growing seasons but during winter, suggesting higher reliability of the model to predict E on a long time scale (e.g., one year).

2. Theory

The model, using the simplified Penman–Monteith equation (McNaughton and Black, 1973; Jarvis and McNaughton, 1986), expresses E as

$$E = \frac{\rho C_p G_c D}{\gamma \lambda} \quad (1)$$

where ρ is the air density, C_p is the specific heat of air, G_c is canopy conductance, D is the vapor pressure deficit, γ is the psychrometric constant, and λ is the latent heat of water vaporization. This equation is derived from the Penman–Monteith equation under the assumption of complete coupling between the canopy and atmosphere.

G_c is formulated as a product of the reference value of G_c when D is 1.0 kPa (G_{cref} , Oren et al., 1999) and functions expressing responses of G_c to meteorological factors (Jarvis, 1976; Lohammer et al., 1980):

$$G_c = G_{cref} \cdot f(D) \cdot g(R) \cdot h(T) \quad (2)$$

where $f(D)$, $g(R)$, and $h(T)$ are functions expressing the responses of G_c to mean daytime D , solar radiation (R), and air temperature (T), respectively. $f(D)$, $g(R)$, and $h(T)$ are respectively modeled as (Oren et al., 1999; Granier et al., 2000b)

$$f(D) = 1.00 - \beta \cdot \ln(D), \quad (3)$$

$$g(R) = \min \left\{ \left(\frac{R}{600} \right)^\delta, 1.00 \right\}, \quad (4)$$

$$h(T) = \begin{cases} 1.00 & (T \geq \epsilon) \\ \frac{T-\zeta}{\epsilon-\zeta} & (\zeta < T < \epsilon), \\ 0.00 & (T \leq \zeta) \end{cases} \quad (5)$$

where β , δ , ϵ , and ζ are parameters. The model does not consider the effect of the soil water deficit on G_c . Most previous studies, making continuous measurements of E (or its substitutes) for forests in Japan including Japanese cedar and cypress plantations (Komatsu et al., 2006; Kosugi et al., 2007; Kumagai et al., 2007), did not report a clear reduction in G_c or E with a soil water deficit, although there were a few exceptions (Tanaka et al., 2002; Komatsu et al., 2007c).

We hypothesized that considering the variation in G_{cref} among sites would be important but considering variations in β , δ , ϵ , and ζ among sites would not be in reproducing E for the sites. Our hypothesis was based on results of previous studies. G_{cref} linearly relates with E in the simplified Penman–Monteith equation. G_{cref} values reported previously (Granier et al., 2000a; Komatsu et al., 2012, 2013) often differ greatly among different forest sites comprising the same species, implying that considering the variation in G_{cref} among the sites is important in estimating E . Oren et al. (1999) and succeeding studies (Addington et al., 2004; Ewers et al., 2008) noted a fairly conservative β among different sites. Thus, assuming β as constant among sites might not introduce large errors in E estimates. Komatsu et al. (2006) analyzed sap flux data on an hourly time scale for a Japanese cedar plantation and

pointed out that the control of G_c by R was important only early in the morning and late afternoon, when E is small. This implies that assuming δ as constant among sites would not introduce large errors in estimating E on a daily time scale. ε and ζ affect E during winter, which accounts for a small portion of annual E for Japanese cedar and cypress plantations (Suzuki, 1980; Kumagai et al., 2014). Thus, assuming ε and ζ as constant among sites might not introduce large errors in estimating E on a longer time scale, such as one year.

We further hypothesized that G_{cref} would relate with the sapwood area on a stand scale (A) for the following reasons. Kumagai et al. (2007) examined E for two cedar forest sites located nearby but having different A . The relative difference in E between the sites was comparable to that in A . Komatsu et al. (2013) compared E values for a cedar forest during two successive growing seasons just before and after thinning. The ratio of E after thinning to that before thinning approximated the ratio of A after thinning to that before thinning. The results of Kumagai et al. (2007) and Komatsu et al. (2013) suggest a correlation between A and G_{cref} . Previous studies (Macfarlane et al., 2010; Ewers et al., 2011) reported relationships between A and E for forests comprising other species. These results also suggest that our hypothesis is reasonable. The relationship between A and G_{cref} would allow the estimation of G_{cref} with the input of A , and then the estimation of E by assuming typical values of β , δ , ε , and ζ and using the simplified Penman–Monteith equation.

3. Data used

We used E data recorded for nine Japanese cedar and five Japanese cypress plantations (Table 1). Twelve of the 14 sites were located in western Japan, but the IK and XT sites were located in eastern Japan and in Taiwan, respectively. Structural parameters differed among the sites. Stem density (N) ranged between 600 and 1575 stems ha^{-1} for cedar and between 350 and 2100 stems ha^{-1} for cypress. The mean diameter at breast height (d_m) ranged between 13.5 and 48.4 cm. The leaf area index (L), estimated using optical methods (see the footnote #2 of Table 1), ranged between 0.8 and 5.9 $\text{m}^2 \text{m}^{-2}$ for cedar and between 0.8 and 4.8 $\text{m}^2 \text{m}^{-2}$ for cypress. Note that L for Japanese cedar and cypress forests estimated using optical methods is generally lower than that estimated using destructive leaf sampling and/or allometry equations (Hasegawa et al., 2013; Tsuruta et al., 2014). The sapwood area at a stand scale (A) ranged between 14.1 and 46.0 $\text{m}^2 \text{ha}^{-1}$ for cedar and between 8.8 and 20.4 $\text{m}^2 \text{m}^{-2}$ for cypress. A was determined by measurements of sapwood thickness using core sampling and assuming the stem cross-section was circular. A complete description of the measurements was provided by Kumagai et al. (2007) and Kume et al. (2010).

E values for all sites were measured employing the sap-flux method and Granier sensors (Granier, 1987). A detailed description of the measurements is available in the papers cited in Table 1.

Table 1

Location, meteorology (the mean annual rainfall P and temperature T), structure (stem density N , the mean diameter at breast height d_m , leaf area index L , and sapwood area A), and description of sap-flux measurements for the plantation sites.

Site ^a	Location	P (mm)	T (°C)	Age (yr)	N (stems ha^{-1})	d_m (cm)	L ($\text{m}^2 \text{m}^{-2}$) ^b	A ($\text{m}^2 \text{ha}^{-1}$)	Plot area (m^2)	Number of the trees in the plot	Period	Reference
<i>Cedar</i>												
IK	36°N, 137°E	2814	12.7	55	600	48.4	3.2	36.6	300	18 (9) ^c	May 15, 2010–May 24, 2011	Shinohara et al. (2013a)
YB	34°N, 131°E	1790	17.2	39	1100	28.9	4.7	22.4	100	11 (11)	Aug 1–Sept 31, 2010	Komatsu et al. (2013)
YA	34°N, 131°E	1790	17.2	40	600	31.6	2.3	14.1	100	6 (6)	Aug 1–Sept 31, 2011	Komatsu et al. (2013)
IS	34°N, 131°E	1807	15.9	43	658	28.2	0.8	21.5	213	14 (11)	May 22, 2012–Feb 28, 2013	Xiang et al. (submitted)
IR	34°N, 131°E	1790	17.2	60	1400	30.3	3.3	28.4	200	28 (14)	July 1–Aug 31, 2010	Ichihashi et al. (in press)
KL	33°N, 131°E	2150	15.0	50	904	40.3	5.7	46.0	321	29 (15)	Mar 3, 2007–July 3, 2008	Kumagai et al. (2007)
KU	33°N, 131°E	2150	15.0	50	1575	23.8	5.4	36.3	318	50 (23)	Mar 3, 2007–July 3, 2008	Kumagai et al. (2007)
KS	33°N, 131°E	2150	15.0	50	1330	26.6	5.9	37.3	203	27 (19)	Mar 3, 2007–July 3, 2008	Kumagai et al. (2014)
XT	24°N, 121°E	2635	17.0	60	625	39.0	2.6	18.4	400	25 (15)	Apr 14, 2012–Mar 19, 2013	Laplace (2013)
<i>Cypress</i>												
IH	34°N, 131°E	1807	15.9	43	863	20.2	0.8	8.8	197	17 (17)	May 22, 2012–Feb 28, 2013	Xiang et al. (submitted)
OL	34°N, 131°E	1790	17.2	49	1450	21.0	3.2	20.4	400	58 (14)	Jan 1–Dec 31, 2008	Kume et al. (2010)
OU	34°N, 131°E	1790	17.2	49	1700	14.9		16.8	100	58 (17)	Jan 1–Dec 31, 2008	Kume et al. (submitted)
SK	34°N, 131°E	1790	17.2	19	2100	13.5	4.8	20.4	100	21 (10)	Apr 1–Aug 15, 2009	Tsuruta et al. (submitted)
HW	34°N, 131°E	1790	17.2	99	350	44.6	3.1	15.2	200	7 (7)	Apr 1–Aug 31, 2009	Tsuruta et al. (submitted)

^a The original names of the sites are given as follows. IK: Ishikawa-ken Forest Experimental Station, YB: Yamanokami site (before thinning), YA: Yamanokami site (after thinning), IS: Yayama Experimental Catchment (cedar) in Iizuka, IR: Ichiripan Plot in the Kasuya Research Forest, KU: UP of Kahoku Experimental Watershed, KL: LP of Kahoku Experimental Watershed, KS: SP of Kahoku Experimental Watershed, XT: Xitou Experimental Forest, IH: Yayama Experimental Catchment (cypress) in Iizuka, OL: Riparian Plot of Ochozu Experimental Watershed, OU: Ridge Plot of Ochozu Experimental Watershed, SK: Sakuta Plot in Kasuya Research Forest, HW: Hiawada Plot in Kasuya Research Forest.

^b L for all sites except OL and IK was measured using a plant-canopy analyzer (LAI-2000, Li-Cor Inc., Lincoln, Nebraska, USA). L for OL and IK was measured using a digital non-spherical color photograph and Gap Light Analyzer software (Franzer et al., 1999).

^c Numerals in the parentheses indicate the number of trees in which sensors were installed.

Employing the method, E was estimated as (Kumagai et al., 2007; Kume et al., 2010)

$$E = \frac{\sum_{i=1}^n F_i \cdot a_i}{S}, \quad (6)$$

where F is the sap flux for an individual tree averaged over its sapwood area, a is the tree sapwood area, n is the number of trees in the plot, and S is the ground area. For a tree whose sapwood thickness was much greater than the sensor length (i.e., 2.0 cm), two or three sensors were installed radially to cover the sapwood area (Kumagai et al., 2007). F was calculated as the weighted average of sap flux for the sensors. Azimuthal variations of sap flux were also considered for some sites (Shinohara et al., 2013a).

4. Methods of analysis

4.1. Determination of parameters

We calculated canopy conductance (G_c) from E estimated using the sap-flux method and meteorological factors for each site using the inverted form of the simplified Penman–Monteith equation:

$$G_c = \frac{\gamma \lambda E}{\rho C_p D}. \quad (7)$$

G_c was calculated as a daily average conductance using mean daytime T and D , and E summed over 24 h but divided by daylight hours (Phillips and Oren, 1998; Kumagai et al., 2008). Sap-flux sensors observe water movement through the trunk during nighttime, which may represent recharge of water into upper sections of the tree trunk and branches or transpiration during nighttime. Dividing E by daylight hours assumes that sap flux observed during nighttime represents recharge of water. Recent studies (Dawson et al., 2007; Fisher et al., 2007; Oishi et al., 2008) have reported that transpiration during nighttime accounts for a considerable portion of daily transpiration. However, transpiration during nighttime would be very low in Japan possibly because of low D during nighttime, as shown by measurements of transpiration using the leaf-chamber method (Kosugi et al., 1995, 1997; Tanaka et al., 2002). The “daytime” here was assumed as the period between 6 a.m. and 6 p.m. throughout the year for simplicity, i.e., the daylight hours were assumed as 12. These assumptions were also made in previous studies (Kumagai et al., 2008; Komatsu et al., 2012). We confirmed that these assumptions did not alter our conclusions. D for the period between 6 a.m. and 6 p.m. differed from D for the period between sunrise and sunset by no more than 8% for our sites. Our preliminary analysis revealed that G_c calculated using D for the period between 6 a.m. and 6 p.m. could differ from those calculated using D for the period between sunset and sunrise by no more than 10%, which was not large enough to alter our conclusions.

We determined G_{cref} , β , and δ for the sites where E data only during the growing season were available. Here, $h(T)$ was assumed to be 1.00 because of relatively high T . We determined G_{cref} , β , δ , ε , and ζ for the sites where year-round E data were available. Note that analysis for the XT site did not follow this policy, as detailed later.

The parameters were determined in a manner similar to that employed by Komatsu et al. (2012, 2013). We first excluded data recorded on rainy days and days with D lower than 0.2 kPa. Data for sap flux (and therefore E and G_c) could suffer from measurement errors on rainy days (Kumagai et al., 2008; Komatsu et al., 2012). G_c could be highly affected by measurement errors in D when D is very low (Ewers and Oren, 2000; Komatsu et al., 2007b). We then examined the relationship between D and G_c using data recorded on days with high R ($>400 \text{ W m}^{-2}$) and T ($>15 \text{ }^\circ\text{C}$). We determined G_{cref} and β by regressing the relationship

employing the least-squares method. After determining G_{cref} and β , we examined the relationship between R and observed G_c divided by $G_{cref} f(D)$ using data recorded on days with high T ($>15 \text{ }^\circ\text{C}$) to determine δ . Here, we confirmed that the relationship between soil water content and observed G_c divided by $G_{cref} f(D)g(R)$ was generally unclear for sites where data for soil water content were available. We finally examined the relationship between T and observed G_c divided by $G_{cref} f(D)g(R)$ to determine ε and ζ . Here, we also confirmed that the relationship between T and observed G_c divided by $G_{cref} f(D)g(R)$ was generally unclear for the sites where E data only during the growing season were available.

We applied a different method to the determination of parameters for XT. XT was located in a mountainous region under a humid subtropical climate. If we had applied the same method as that applied to other sites, we would have only limited data with which to examine the relationship between D and G_c because of a large number of rainy days and days with low R . The method applied to XT was as follows. We first excluded data recorded on days with daily rainfall more than 5.0 mm and days with D lower than 0.2 kPa. We then examined the relationship between D and G_c using data recorded on days with R no less than 300 W m^{-2} . We determined G_{cref} and β by regressing the relationship employing the least-squares method. Here, we confirmed that there was no systematic difference in G_c between days without rain and days with rain no more than 5.0 mm. After determining G_{cref} and β , we examined the relationship between R and observed G_c divided by $G_{cref} f(D)$ to determine δ . In the above analysis, we did not exclude data using a criterion about T , because the range of T was narrow for XT and a large portion of the data satisfied T being no less than $15 \text{ }^\circ\text{C}$. We did not determine the parameters in $h(T)$, because we did not find a clear relationship between T and observed G_c divided by $G_{cref} f(D)g(R)$ for XT.

4.2. Assessing the importance of parameters

We hypothesized that considering the variation in G_{cref} among the sites would be important but considering variations in β , δ , ε , and ζ among the sites would not be in reproducing E for the 14 sites (detailed in the theory). To verify this hypothesis, we calculated E for days with no rain and T no less than $15 \text{ }^\circ\text{C}$ (corresponding to a growing season) for the sites using three different parameterizations: (P1G) the mean values of G_{cref} , β , and δ among the sites, (P2G) G_{cref} determined for each site and the mean values of β and δ among the sites, and (P3G) G_{cref} , β , and δ determined for each site. To evaluate the reproducibility of E by P1G, P2G, and P3G, we examined the relationships of the mean E for the days calculated using P1G, P2G, and P3G respectively with the mean observed E . If the difference in reproducibility between P1G and P2G was found to be important, considering the variation in G_{cref} among the sites would be important in calculating the mean E for a growing season. If the difference in reproducibility between P2G and P3G was found to be unimportant, considering variations in β and δ would not be important.

Similarly, we calculated E for days with no rain for sites having data recorded in winter as well as those recorded in a growing season using three different parameterizations: (P1W) the mean values of G_{cref} , β , δ , ε , and ζ among the sites, (P2W) G_{cref} determined for each site and the mean values of β , δ , ε , and ζ among the sites, and (P3W) G_{cref} , β , δ , ε , and ζ determined for each site. To evaluate the reproducibility of E by P1W, P2W, and P3W, we examined the relationships of the mean E for the days calculated using P1W, P2W, and P3W respectively with the mean observed E . If the difference in reproducibility between P1W and P2W was found to be important, considering the variation in G_{cref} would be important in calculating the mean E for a period including a growing season and winter. If the difference in reproducibility between P2W and

P3W was found to be unimportant, considering variations in β , δ , ε , and ζ among the sites would not be important.

When evaluating the model reproducibility, we primarily focused on the mean E on a longer time scale, because our model was intended to be used to improve our understanding of the forest water balance on a long time scale (see Section 1). For assessing the importance of the difference in reproducibility between different parameterizations, we did not examine whether there was a statistically significant difference between errors in E estimates using different parameterizations. What we need to know is not whether there is a statistical difference, but whether the magnitude of the difference is important in a practical context (Bakan, 1966; Thompson, 1996, 1998; Nuzzo, 2014). We thus examined whether errors in estimates of the mean E made using the G_c model were less than potential errors in observed E . Kumagai et al. (2007) and Kume et al. (2010) examined uncertainty in E estimates made using the sap-flux method for Japanese cedar and cypress plantations. Sampling data for sap flux on a sensor scale repeatedly using the Monte-Carlo technique, Kumagai et al. (2007) and Kume et al. (2010) revealed that the standard deviation of E estimates was generally 10% or more of the mean E . When considering that 95% of data theoretically fall in the range of two standard deviations above and below the mean, potential errors in observed E would be less than 20% of the value in most cases. If errors in E estimates made using different parameterizations were more than 20% of observed E , we would regard the difference in reproducibility as important.

In addition, we examined whether errors in estimates of the mean E made using the G_c model were less than those in interception evaporation estimates made using the model developed by Komatsu et al. (2007a) or its revised form developed by Komatsu et al. (submitted), because our G_c model was intended to be used with the interception evaporation model to assess the water balance for Japanese cedar and cypress plantations (Komatsu et al., submitted). The typical errors in interception evaporation estimates for the period with T no less than 15 °C and for a year would be 0.22 and 0.20 mm day⁻¹, respectively (Appendix A).

4.3. Relating important parameters with species and structures

We examined the relationship between A and G_{cref} to model G_{cref} using A . After confirming correlation between the two variables, we regressed the relationship to obtain a linear equation relating G_{cref} with A . The intercept of the equation was assumed to be zero. The slope of the equation was determined employing the least-squares method. To examine the stability of the correlation and the slope, we calculated 95% confidence intervals of these variables employing the bootstrapping method (Efron, 1979; Diadonis and Efron, 1983). These intervals were “bias-corrected, accelerated” percentile intervals calculated in the manner described by Efron (1987) and Fox (2008).

As data for A are not usually available for most Japanese cedar and cypress plantations, we tried to relate A with the mean diameter at breast height (d_m) and stem density (N) data in the following way. Tsuruta et al. (2011) summarized data for the sapwood area at a tree scale (a) for 81 cedar trees from six sites and 109 cypress trees from nine sites and examined the relationships between diameter at breast height (d) and a for these species. They did not find clear differences in the relationship among sites. They regressed the data to develop general equations for the relationships: $a = a_{ref} \cdot d^k$, where a and d were respectively in units of cm² and cm and a_{ref} and k were parameters. a_{ref} was 1.40 and 1.96 cm² for cedar and cypress, respectively. k was 1.55 and 1.42 for cedar and cypress, respectively. Approximating this equation by the tangential line at d being equal to d_m , A was expressed as

$$A = N \cdot a_{ref} \cdot d_m^k \tag{8}$$

To examine the validity of this method of estimating A from N and d_m , we investigated the relationship between A estimated employing this method and observed A . For this validation, we used A data for YB, YA, IS, KL, KU, KS, XT, and IH, and another cypress site of Sun et al. (2014). Data for the other six sites (i.e., IK, IR, OL, OU, SK, and HW) were not used, because a data for the sites were used by Tsuruta et al. (2011) to develop the relationship between d and a .

5. Results

5.1. Determination of parameters

Fig. 1 shows the relationships between D and G_c , between R and G_c divided by $G_{cref}f(D)$, and between T and G_c divided by $G_{cref}f(D)$

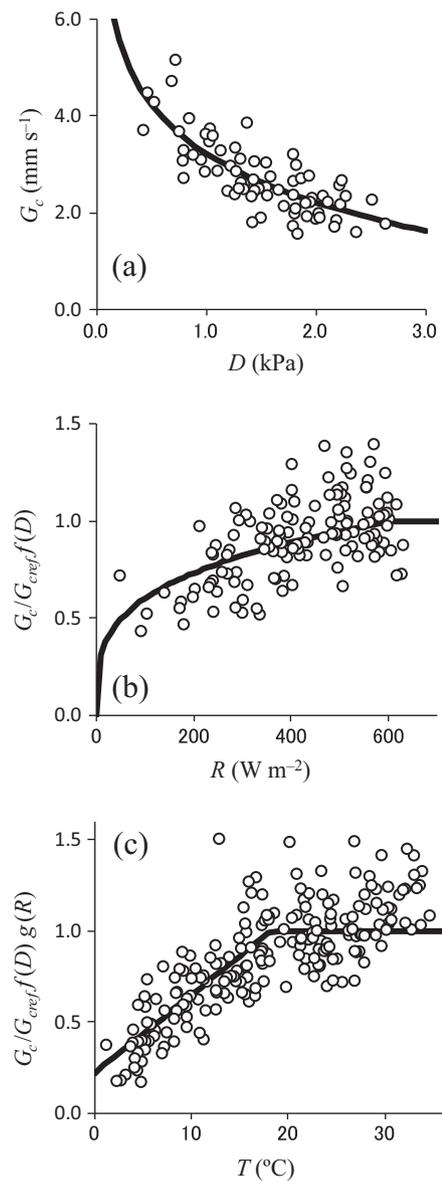


Fig. 1. Relationships (a) between vapor pressure deficit (D) and canopy conductance (G_c), (b) between solar radiation (R) and G_c divided by $G_{cref}f(D)$, and (c) between temperature (T) and G_c divided by $G_{cref}f(D)g(R)$ for the OL site. Solid lines in (a–c) indicate $f(D)$, $g(R)$, and $h(T)$ for OL, respectively. G_{cref} is the reference value of canopy conductance. $f(D)$, $g(R)$, and $h(T)$ are functions expressing the responses of G_c to D , R , and T .

Table 2
Parameter values optimized for each site.

Site	G_{cref} (mm s ⁻¹)	β in $f(D)$	R^2	δ in $g(R)$	R^2	ε (°C) in $h(T)$	ζ (°C) in $h(T)$	R^2
<i>Cedar</i>								
IK	9.77	0.750	0.782	0.287	0.389	7.75	0.313	0.469
YB	3.46	0.441	0.493	0.0893 ^a	0.00793			
YA	1.89	0.427	0.365	0.0625 ^a	0.00247			
IS	2.84	0.569	0.580	0.238	0.339	20.7	-48.8	0.164
IR	2.81	0.351	0.214	0.288	0.219			
KL	6.07	0.783	0.605	0.347	0.316	12.3	0.0649	0.587
KU	4.22	0.749	0.438	0.410	0.240	15.1	-5.22	0.343
KS	5.19	0.667	0.628	0.378	0.392	14.8	-2.49	0.544
XT	2.84	0.412	0.213	0.308	0.310			
<i>Cypress</i>								
IH	1.76	0.472	0.622	0.000 ^a	-0.0329	21.0	-2.84	0.740
OL	3.22	0.453	0.628	0.285	0.330	18.2	-5.19	0.636
OU	1.86	0.591	0.738	0.417	0.561	17.2	-5.78	0.584
SK	5.65	0.468	0.412	0.432	0.252			
HW	2.81	0.688	0.436	0.441	0.270			
Mean	3.89	0.556		0.285		15.9	-8.74	
Standard deviation	2.19	0.146		0.142		4.43	16.4	

^a A small δ value does not suggest that there was no effect of solar radiation on transpiration, but suggests that the effect was not detectable by an analysis on a daily time scale.

$g(R)$. Here, we show data only for OL as an example. Regressing these relationships and relationships for the other sites, parameters for the sites were determined as listed in Table 2. The maximum G_{cref} (9.77 mm s⁻¹ for IK) among the sites was more than five times the minimum G_{cref} (1.76 mm s⁻¹ for IH). We observed variations in $f(D)$, $g(R)$, and $h(T)$ among sites (Fig. 2).

5.2. Assessing the importance of parameters

Fig. 3 shows the relationships of the mean E for days with no rain and T no less than 15 °C calculated using P1G, P2G, and P3G with observed E . E calculated using the model was not strongly correlated ($r = 0.454$) with observed E for P1G, but was strongly correlated ($r = 0.993$ and 0.997 , respectively) and fell along the 1:1 line for P2G and P3G. Errors for P1G were often greater than potential errors in observed E and interception evaporation estimates, but errors for P2G and P3G were not. Fig. 4 shows the relationships of mean E for days with no rain calculated using P1W, P2W, and P3W with observed E for sites having data recorded in winter as well as those recorded in a growing season. E calculated using the model was not strongly correlated ($r = 0.597$) with observed E for P1W, but was strongly correlated ($r = 0.999$ and >0.999 , respectively) and fell along the 1:1 line for P2W and P3W. Errors for P1W were often greater than potential errors in observed E and interception evaporation estimates, but errors for P2W and P3W were not. These results suggest that considering the variation in G_{cref} among the sites was important for reproducing the mean E on a long time scale, but considering variations in β , δ , ε , and ζ was not.

Besides the mean E on a long time scale, we also examined model reproducibility of seasonal and day-to-day variations in E using data on a daily time scale for sites where year-round data were available. Fig. 5 shows time series of E calculated using P1W, P2W, and P3W and observed E . Here, we show data only for OU as an example. The slope of the regression line for the relationship between calculated and observed E often fell outside the range between 0.8 and 1.2 for P1W, where the range corresponded to the uncertainty in observed E . However, the slope always fell in the range for P2W and P3W (Table 3). The determination coefficients for P1W and P2W were often lower than that for P3W. However, the difference was relatively small. These results suggest that considering the variation in G_{cref} among the sites was important for

reproducing seasonal and day-to-day variations in E , but considering variations in β , δ , ε , and ζ was not.

5.3. Relating important parameters with species and structures

G_{cref} tended to increase with A (Fig. 6). The correlation coefficient (r) was 0.698 and the 95% confidence interval was (0.446, 0.951), when pooling data for cedar and cypress. The correlation between A and G_{cref} was stronger than that between L and G_{cref} ($r = 0.479$). The regression line was determined as $G_{cref} = 0.157 A$, where A and G_{cref} were in units of m² ha⁻¹ and m s⁻¹, respectively. The 95% confidence interval of the slope was (0.127, 0.195). Data for both cedar and cypress were located along the regression line, suggesting no clear difference in G_{cref} for a given A between cedar and cypress. Thus, the regression equation could be used to predict G_{cref} from A for both cedar and cypress plantations and then to predict E . Fig. 7 shows the normal probability plot (Fujii, 2005; Peck and Devore, 2005) for differences between G_{cref} estimated using the regression equation and observed G_{cref} , where the normal score in the y-axis indicates the difference divided by the standard deviation of the difference for the 14 sites. Data for all sites except SK and IK were approximated by a line, indicating that most of these data followed a normal distribution. However, data for SK and IK were located far from the line, suggesting that observed G_{cref} for these sites was higher than expected from the regression equation and a normal distribution of the differences between estimated and observed G_{cref} .

Fig. 8 compares A estimated from N and d_m and observed A . Data were generally located around the 1:1 line. The mean relative error (i.e., the ratio of the difference between estimated and observed A to observed A) was 26%. The magnitude of this error is discussed in Section 6.3.

6. Discussion

6.1. Roles of $f(D)$, $g(R)$, and $h(T)$

Our model succeeded in reproducing E observed for the 14 sites, suggesting validity of our hypothesis. Our model includes three functions, i.e., $f(D)$, $g(R)$, and $h(T)$. These functions influence on calculated E differently. Omitting $f(D)$ from Eq. (2) (i.e., assuming $f(D)$ being 1.0) causes overestimation of E during May–August (Fig. 9a),

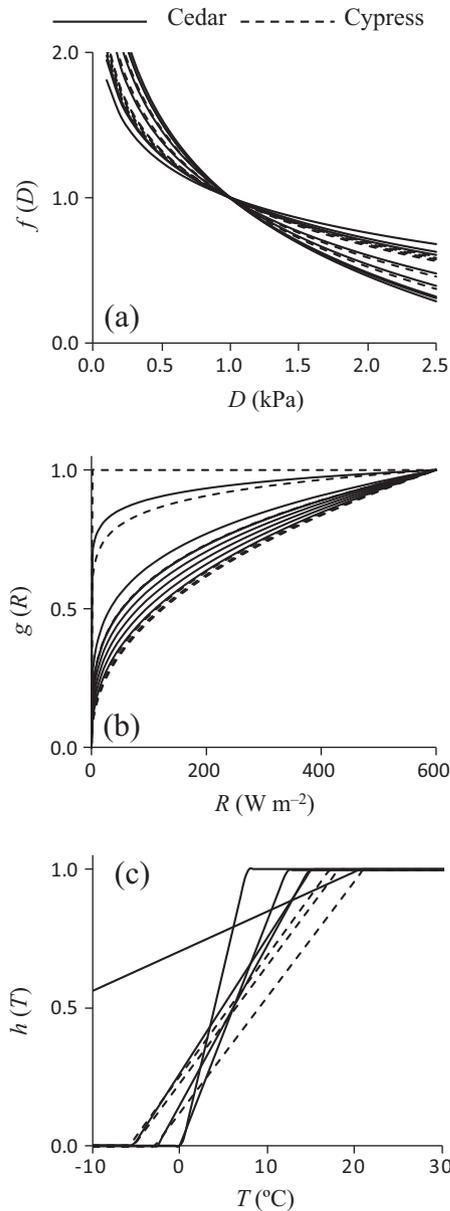


Fig. 2. Functions expressing the responses of canopy conductance to (a) vapor pressure deficit ($f(D)$), (b) solar radiation ($g(R)$), and (c) temperature ($h(T)$) for each site.

resulting in a relatively low determination coefficient ($R^2 = 0.740$) for the relationship between calculated and observed E . This suggests that considering reduction in G_c with increasing D is important in reproducing E when D is high.

Omitting $g(R)$ causes overestimation of E during October–January when R is relatively low. Omitting $h(T)$ causes overestimation of E during January–February when T is relatively low. However, the determination coefficients for these cases ($R^2 = 0.792$ and 0.798 , respectively) are higher than that for the case of omitting $f(D)$. This suggests that effects of omitting $g(R)$ and $h(T)$ on reproducing E on a daily time scale are less important than that of omitting $f(D)$.

6.2. Variability in G_{cref}

G_{cref} values for SK and IK were higher than expected from the regression equation (Fig. 7). There are technical factors that

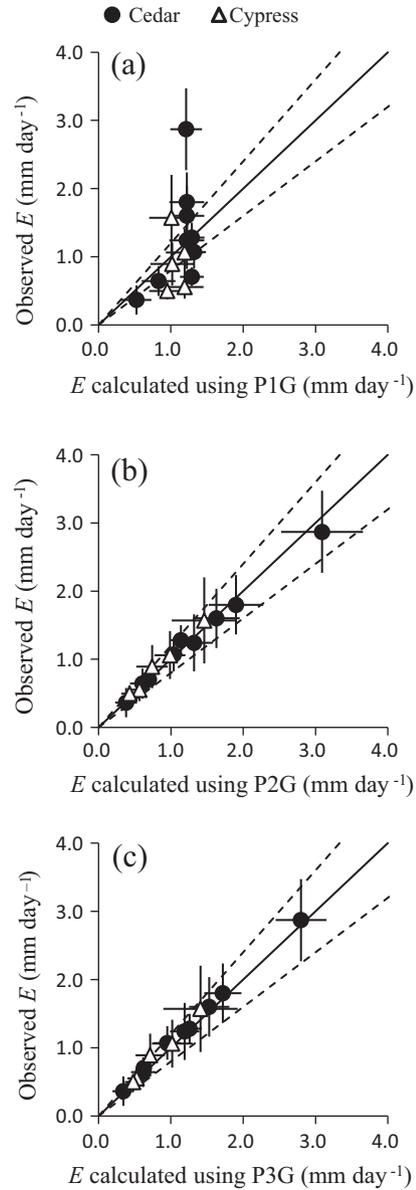


Fig. 3. Relationships of the mean transpiration (E) for days with no rain and temperature being no less than $15\text{ }^\circ\text{C}$ (i.e., corresponding to a growing season) calculated using (a) P1G, (b) P2G, and (c) P3G with the mean observed E . Here, P1G, P2G, and P3G respectively use the mean values of G_{cref} , β , and δ among the sites, G_{cref} determined for each site and the mean values of β and δ among the sites, and G_{cref} , β , and δ determined for each site. The solid line indicates the 1:1 relationship. Dotted lines indicate conditions that errors in E estimates equal potential errors in observed E . Error bars indicate standard deviations.

potentially affect variations in G_{cref} among sites. An insufficient number of sensors for sap flux measurements could result in large errors in E estimates. Shinohara et al. (2013a) compared errors in E estimates introduced by ignoring tree-to-tree variations, radial variations, and circumference variations in sap flux for a Japanese cedar plantation. They concluded that tree-to-tree variations in sap flux are the primary factor of errors in E estimates. Previous studies (Kumagai et al., 2005, 2007; Kume et al., 2010; Shinohara et al., 2013b) reported that errors in E estimates for Japanese cedar and cypress were serious when the number of trees in which sensors were installed was low, especially when there were less than approximately 10. Note that this threshold number would be species specific, because different threshold numbers were reported for other species (Oren et al., 1998; Mackay et al., 2010). There

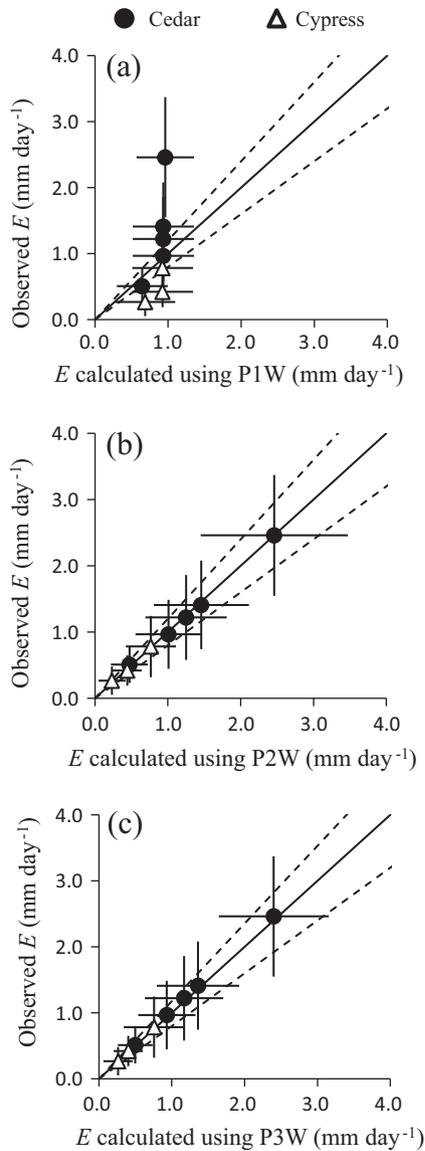


Fig. 4. Relationships of the mean transpiration (E) for days with no rain calculated using (a) P1W, (b) P2W, and (c) P3W with the mean observed E for sites having data recorded in winter as well as those recorded in a growing season. Here, P1W, P2W, and P3W respectively use the mean values of G_{cref} , β , δ , ϵ , and ζ among the sites, G_{cref} determined for each site and the mean values of β , δ , ϵ , and ζ among the sites, and G_{cref} , β , δ , ϵ , and ζ determined for each site. The solid line indicates the 1:1 relationship. Dotted lines indicate conditions that errors in E estimates equal potential errors in observed E . Error bars indicate standard deviations.

were approximately 10 in which sensors were installed at each site (10 for SK and 9 for IK). Furthermore, E for IK estimated from sap flux data derived for nine trees differed less than 5% from that estimated from sap flux data derived for all 18 trees according to intensive measurements performed during May 16–18 and June 1–4, 2010 for the site. Thus, the numbers of sensors at these sites are not likely a main factor explaining the higher G_{cref} .

The use of the simplified Penman–Monteith equation could also be a factor causing higher G_{cref} . Aerodynamic conductance, which is assumed as infinite in the simplified Penman–Monteith equation, generally ranges between 70 and 400 mm s^{-1} for coniferous forests including Japanese cedar and cypress plantations (Stewart and Thom, 1973; Yamanoi and Ohtani, 1992; Loustau et al., 1996; Tanaka et al., 1996). When assuming aerodynamic conductance as 70 mm s^{-1} and calculating G_{cref} using the original Penman–Monteith equation (see the methods used by Komatsu et al.,

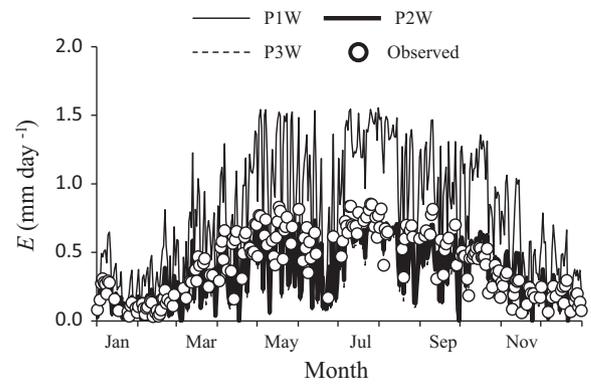


Fig. 5. Time series of daily transpiration (E) calculated using P1W, P2W, and P3W and observed E for OU. Lines for P2W and P3W overlap each other. E observed on rainy days is not plotted in this figure.

Table 3
Slope and coefficient of determination (R^2) for the relationship between calculated and observed transpiration.

Site	P1W		P2W		P3W	
	Slope ^a	R^2	Slope	R^{2b}	Slope	R^2
<i>Cedar</i>						
IK	2.43	0.329 ^c	0.946	0.329	1.02	0.576
IS	0.776	0.867	1.06	0.867	1.03	0.874
KL	1.50	0.845	0.964	0.845	1.04	0.818
KU	1.05	0.757	0.970	0.757	1.06	0.726
KS	1.32	0.848	0.988	0.848	1.05	0.845
<i>Cypress</i>						
IH	0.510	0.927	1.13	0.927	1.02	0.937
OL	0.879	0.843	1.06	0.843	1.04	0.883
OU	0.466	0.891	0.974	0.891	1.05	0.911

^a The intercept of the regression equation was assumed to be zero.
^b Determination coefficients for P2W are identical to those for P1W, because the responses of canopy conductance to meteorological factors for P2W are same as those for P1W.
^c Low determination coefficients for IK were primarily caused by the small number of data with low E due to rejection of data recorded on rainy days. When including data recorded on rainy days in the analysis, the determination coefficients were 0.829, 0.829, and 0.858 for P1W, P2W, and P3W, respectively.

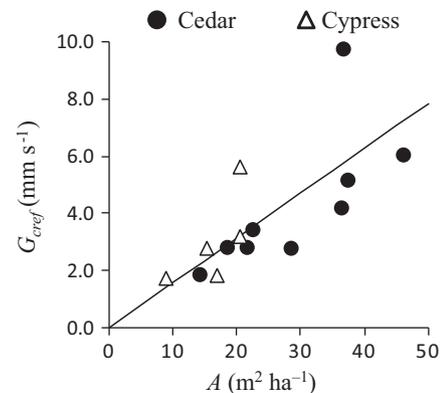


Fig. 6. Relationship between the sapwood area at a stand scale (A) and the reference value of canopy conductance (G_{cref}). The regression line for the relationship between A and G_{cref} , determined using the least-squares method, is written as $y = 0.157x$.

2012), G_{cref} values for SK and IK are determined as 3.53 and 6.22 mm s^{-1} , respectively. These G_{cref} data are located close to the regression line in Fig. 6, implying that the use of the simplified Penman–Monteith equation might be a potential factor explaining the higher G_{cref} for the sites. Unfortunately, there have been no data

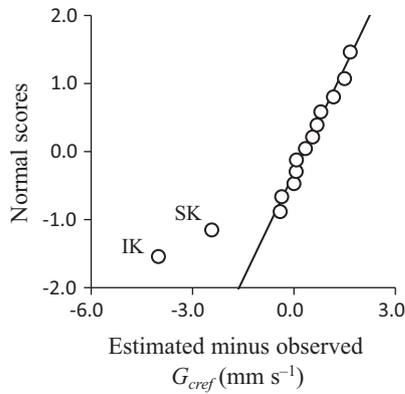


Fig. 7. Normal probability plot for the difference between estimated and observed reference values of canopy conductance (G_{cref}). The solid line is the regression line, determined using the least-squares method, for all data except those for SK and IK.

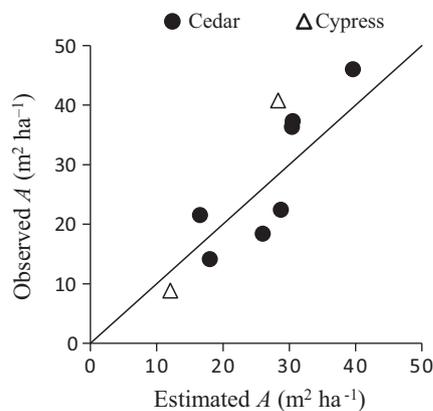


Fig. 8. Comparison between estimated and observed sapwood areas on a stand scale (A). The solid line indicates the 1:1 relationship.

for aerodynamic conductance observed at these sites. However, data for wind speed for IK are available. The relationship between D and G_c differed only slightly according to wind speed, suggesting that low aerodynamic conductance would not be the primary factor causing low G_{cref} for IK. Data for wind speed for SK were unavailable.

Besides the factors discussed above, age might be another possible factor explaining high G_{cref} for SK, which is younger than the other sites (Table 1). There have been several studies reporting or suggesting changes in sap flux on a stand scale with age for mono-specific forests (Tsuruta et al., 2008; Forrester et al., 2010). We do not have data for other sites of similar age to SK. We thus recommend examining G_{cref} for young Japanese cedar and cypress plantations.

6.3. Errors in A estimates

We observed a clear correlation between estimated and observed A (Fig. 8). However, the error in A estimates was relatively large. The mean relative error in A estimates (26%) was comparable to that in G_{cref} estimates obtained using the regression equation in Fig. 6 (24%). Thus, it would be better to use observed A to calculate G_{cref} , if observed A is available.

The error in A estimates would be primarily caused by errors in a estimates using Tsuruta et al.'s (2011) equation, and only secondarily caused by errors due to tangential approximation of the equation (Eq. (8)). The tangential approximation could cause underestimation of A , because a values calculated using the

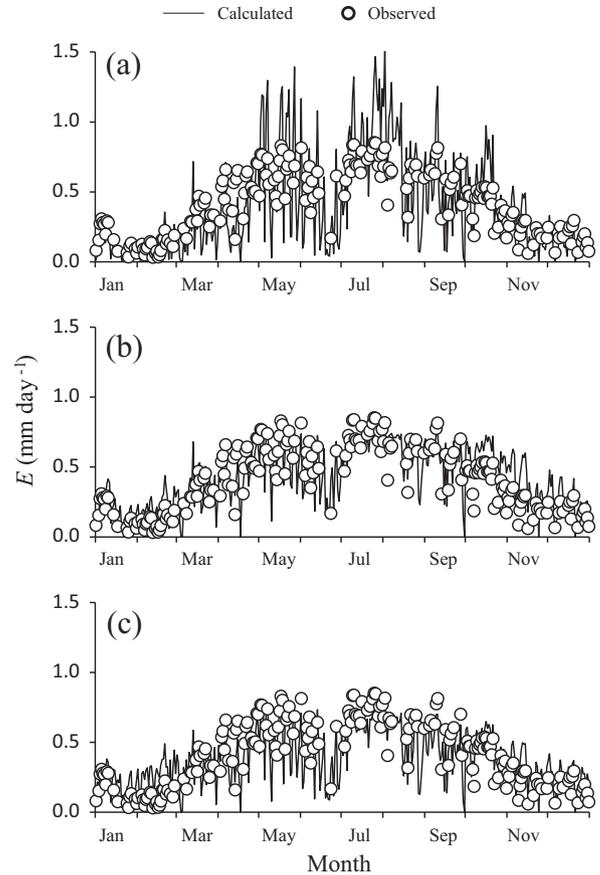


Fig. 9. Time series of daily transpiration (E) calculated using P2W by omitting functions expressing the responses of G_c to (a) vapor pressure deficit ($f(D)$), (b) solar radiation ($g(R)$), and (c) temperature ($h(T)$).

tangential line are no more than those calculated using the original equation. In fact, A estimated employing the above method did not always underestimate observed A (Fig. 8).

6.4. Possible applications and implications

Our model would be useful as a research tool for hydrologists. There have been many studies (Dung et al., 2012; Kubota et al., 2013) examining changes in the annual water balance with changing structure of Japanese cedar and cypress plantations due to forest management. There are large variations in the change in the annual water balance among studies. Our model, accompanied with the interception evaporation model developed by Komatsu et al. (2007a), could be used by hydrologists to calculate changes in E and interception evaporation for catchments and improve our understanding of underlying processes of the variations. Our model and the interception model are also useful in estimating spatial variations in E and interception evaporation on a landscape scale, when these models are used with forest inventory data.

As described above, our model has great potential for application. This is because our model can be used only with the inputs of commonly available data for forest structure (i.e., N and d_m) and meteorology (see Appendix B). Our model is specific to Japanese cedar and cypress plantations. However, our model is important in that it demonstrates how to relate E with commonly available data. The concept of our model would be useful in developing similar models to estimate E for mono-specific forests comprising other species.

On the other hand, the model needs to be tested further. The sites for cedar and cypress plantations (Table 1) were located

mainly in western Japan, where temperature is intermediate or higher in the distribution areas of cedar and cypress (Japan Forestry Agency, 2013). We do not have enough E data for cedar and cypress plantations recorded in regions where temperature is lower, except data for IK. The response of G_c to temperature for stands in this region might be different from that observed in this study. We thus recommend testing the applicability of the model using data derived for sites located in this region. Furthermore, it is preferable to test the model applicability using data for sites with various age classes. Ages of the sites (Table 1) ranged between 19 and 99 years, which almost covers the age range for most Japanese cedar and cypress plantations in Japan (Japan Forestry Agency, 2013). However, many of the sites fall in the range of 40–60 years. Thus, applicability of our model has been tested sufficiently for stands within this age class, but has not been fully tested outside the age class. A major portion of Japanese cedar and cypress plantations in Japan falls in this age class, suggesting the practical usefulness of the model. On the other hand, the portion of older stands is expected to increase, because forestry in Japan has stagnated and cedar and cypress plantations have not been actively harvested (Komatsu et al., 2010). Therefore, we recommend testing the applicability of the model to stands outside the age class.

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Appendix A. Errors in interception evaporation estimates

The interception evaporation model developed by Komatsu et al. (2007a) and its revised form developed by Komatsu et al. (submitted) typically have an error of 4% of incident rainfall. The period with T no less than 15 °C is typically April–October in regions where Japanese cedar and cypress are distributed and incident rainfall during the period is typically 1200 mm (National Astronomical Observatory, 2013). Annual rainfall in the regions is typically 1800 mm. Thus, the typical error in interception evaporation estimates for the period with T no less than 15 °C would be 48 mm, which was equivalent to 0.22 mm day⁻¹. The typical error for a year would be 72n the basis of observed Amm, which was equivalent to 0.20 mm day⁻¹.

Appendix B. Methods of preparing meteorological inputs

The model developed in this study requires daytime (6 a.m.–6 p.m.) mean solar radiation (R), temperature (T), and vapor pressure

deficit (D) as meteorological inputs. R data are not always available throughout Japan (Shinohara et al., 2007). It is often the case that only daily maximum (T_x) and minimum temperatures (T_n) are available for historical data (The University of Tokyo Forests, 2014).

The former problem can be solved by substituting R for the target area by data for solar radiation recorded at a meteorological observatory in surrounding areas. Data for daily solar radiation are recorded at main meteorological observatories in Japan (Japan Meteorological Agency, 2014). R in units of $W m^{-2}$ can be approximated by daily solar radiation in units of $MJ m^{-2} day^{-1}$ multiplied by an index of 23.1, which converts the units. E during January 1–December 31, 2008 calculated using the model with the input of R , T , and D observed at OL is 234 mm. Here, G_{cref} is assumed to be 3.22 mm s⁻¹ on the basis of observed A and the relationship between A and G_{cref} (Fig. 6). E values during the same period calculated with the input of solar radiation observed at meteorological observatories at Fukuoka, Saga, Oita, and Hiroshima are 238, 234, 237, and 240 mm, respectively. These meteorological observatories are located 15, 50, 110, 200 km from OL, respectively. All these values approximate the value calculated using R for OL. Qualitatively, the same results are available for other sites. Thus, accurate estimates of R are not important for estimating E using the model. This is attributed to E being insensitive to R when R is high (Fig. 2b) and E on days with high R accounting for a relatively large portion of annual E .

The latter problem (i.e., only T_x and T_n are available) can be solved making the following assumptions. First, the diurnal trend in temperature is approximated using a sine function minimized at 6 a.m. and maximized at 2 p.m. A similar approximation has been commonly used to produce hourly temperature data from T_x and T_n (Campbell and Norman, 1998). Under this assumption, T (i.e., temperature during 6 a.m.–6 p.m.) is analytically written as $T = (T_x + T_n)/2 + (T_x - T_n)/(3\pi)$. Second, vapor pressure deficit at 6 a.m. is zero and daytime vapor pressure deficit is caused by a temperature rise during the day. Note that the assumption of vapor pressure deficit being zero in the morning is generally valid except in arid and semi-arid regions (Running et al., 1987; Kimball et al., 1994). D (i.e., the mean vapor pressure deficit during 6 a.m.–6 p.m.) is thus approximated by $D = e_s(T) - e_s(T_n)$, where e_s is the saturation vapor pressure. E during January 1–December 31, 2008 calculated using the model with the input of R , T , and D observed at OL is 234 mm. E during the same period calculated with the input of R , T_x and T_n observed at OL is 225 mm, which approximates E calculated with the input of T and D . Qualitatively, the same results are available for other sites.

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